TESTING IN NONPARAMETRIC ACCELERATED LIFE TIME MODELS

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We consider a random life time T which depends on some explanatory variable or covariate X in such way that X reduces a basic life time, say T_0 , by a factor $\psi(X)$. In other words, the life time T satisfies the equation

$$T = \frac{T_0}{\psi(X)}.$$

Examples for X are the dose of a drug, temperature, pressure or stress. The function ψ is called *acceleration function*. We will assume that T is an absolute continuous random variable and that the covariate X does not depend on the time; and for simplicity of presentation let X take one dimensional values.

The conditional survival function of T given X = x has the form

$$S(t|x) = \mathsf{P}(T > t|X = x) = S_0(t\psi(x))$$

where S_0 is the survival function of T_0 , the so-called baseline survival function.

The aim of the talk is to propose a test procedure for testing whether the function ψ belongs to a pre-specified parametric class of functions

$$\mathcal{F} = \{\psi \,|\, \psi(\cdot) \,=\, \psi(\cdot; eta) \;eta \in \mathbb{R}^d\}.$$

The test is based on the following transformation: The conditional expectation of $Y = \log T$ is given by

$$\mathsf{E}(Y|X=x) = \mu_0 - \log \psi(x)$$
 with $\mu_0 = -\int \log z \, \mathrm{d}S_0(z) = \mathsf{E}(\log T_0),$

and we can translate the considered test problem into a problem of testing the regression function in a nonparametric regression model

$$Y = m(X) + \varepsilon,$$

where

$$m(x) = \mu_0 - \log \psi(x)$$
 and $\mathsf{E}(\varepsilon|x) = 0.$

Now, define the set of hypothetical regression functions

$$\mathcal{M} = \{ m(\cdot; \vartheta) = \mu_0 - \log \psi(\cdot; \beta), \ \vartheta = (\mu_0, \beta) \in \mathbb{R}^{d+1} \}.$$

Then the "new" test problem has the form

$$H: m \in \mathcal{M}$$
 versus $K: m \notin \mathcal{M}$.

Before we will give the test statistic let us describe the underlying data: We do not observe the pairs (X_i, T_i) but consider a model in which the life times T_i are subject to random right censoring, so that the observable random variables are given by

$$(X_i, T_i^*, \Delta_i)$$
 with $T_i^* = \min(T_i, C_i)$ and $\Delta_i = 1(T_i \leq C_i), \quad i = 1, \dots, n$

Here C_i is a nonnegative random variable, representing the censoring time. To identify the conditional distribution of T given X we will assume that C and (X, T) are independent, or at least that T and C are conditionally independent given X.

The test statistic is based on the following idea: Since the alternative consists of "all" regression functions we compare a nonparametric estimator of m, say \hat{m}_n , with an estimator for the hypothetical parametric $m(\cdot; \vartheta)$, i.e. $m(\cdot; \hat{\vartheta}_n)$. But since the estimate \hat{m}_n is a result of smoothing it seems to be appropriate to compare \hat{m}_n with a smoothed version of $m(\cdot; \hat{\vartheta}_n)$, denoted by $\tilde{m}_n(\cdot; \hat{\vartheta}_n)$. As deviation measure we take the weighted integrated squared distance. Proving the asymptotic normality of this (properly standardized) statistic we obtain an asymptotic α -test.

This idea is realized as follows: Since

$$m(x) = -\int_0^\infty \log t \, \mathrm{d}S(t|x)$$

a nonparametric estimator of m is given by

$$m(x) = -\int_0^\infty \log t \, \mathrm{d}\hat{S}_n(t|x)$$

where \hat{S}_n is the weighted Kaplan-Meier estimator for the conditional survival function. As weights, denoted by W_{ni} , we take kernel weights. The Kaplan-Meier estimator is the extension of the empirical survival function for censored data, and the weighted Kaplan-Meier estimator is the analog for conditional survival functions.

Further, let $\hat{\vartheta}_n$ be the least squares estimator of the unknown parameter in the hypothetical regression model, and set

$$\tilde{m}_n(x;\hat{\vartheta}_n) = \sum_i W_{ni}(x)m(X_i,\hat{\vartheta}_n).$$

The test statistic is defined by

$$Q_n = \int [\hat{m}_n(x) - \tilde{m}_n(x; \hat{\vartheta}_n)]^2 a(x) \, \mathrm{d}x,$$

here a is a known weight function which is introduced to control the region of integration. Using standard methods for deriving asymptotic normality of quadratic forms we obtain

$$\rho_n \left(Q_n - e_n \right) \xrightarrow{\mathcal{D}} \mathcal{N}(0, 1)$$

for standardizing terms ρ_n and e_n , which depend on the underlying unknown distribution. Replacing these unknown terms by appropriate estimates $\hat{\rho}_n$ and \hat{e}_n , respectively, we get the desired asymptotic α -test: Reject H if

$$Q_n > z_\alpha \hat{\rho}_n + \hat{e}_n,$$

where z_{α} is the fractile of the standard normal distribution.

The proposed procedure will be completed by simulation results concerning this "limit-distribution-approach" and some resampling methods for the test statistic Q_n .